



THE UNIVERSITY *of* EDINBURGH
informatics

Applied Machine Learning (AML)

Introduction to Classification

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Classification

Classification Overview

- In supervised learning, we are tasked with predicting an output y , given an input feature vector x
- For **classification** problems, the output space is a set of mutually exclusive ‘classes’ (also commonly referred to as ‘labels’)

Binary versus Multiclass Classification

- In **binary classification** we have two possibilities, e.g. dog versus cat. Thus, $y \in \{0, 1\}$, $y \in \{1, 2\}$, $y \in \{-1, +1\}$, ...
- In **multiclass classification** we can have C possible options, e.g. different breeds of dog. Thus, $y \in \{1, \dots, C\}$, where C is the number of classes of interest

Example Classification Problems

- Spam filtering
- Determining the object present in an image, i.e. image classification
- Fraudulent transaction detection
- Music genre classification
- Medical diagnostic tests
- ...

Example 1D Classification Problem

- We have collected a dataset containing the measurements of the petal lengths (in cm) of plants from two different species: **species A** and **species B**
- Thus, we have a one dimensional (1D) continuous measurement $x \in \mathbb{R}$ and a binary class label $y \in \{0, 1\}$



Example 1D Classification Problem

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- Thus, we have a one dimensional (1D) continuous measurement $x \in \mathbb{R}$ and a binary class label $y \in \{0, 1\}$
- For species A, we have five measurements $\{1.8, 2.1, 2.5, 3.2, 3.8\}$ and for species B we have three $\{5.8, 6.7, 7.0\}$
- We can write our dataset $\mathcal{D} = \{(x_n, y_n)\}_{n=1}^N = \{(1.8, 0), (2.1, 0), (2.5, 0), (3.2, 0), (3.8, 0), (5.8, 1), (6.7, 1), (7.0, 1)\}$



The Generative Approach

- Given a new observation x , can we predict which of the two classes it most likely belongs to?
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- We can then use this model to make predictions about unobserved (i.e. *new*) data
- For *continuous* features, one obvious choice is the **Gaussian** distribution

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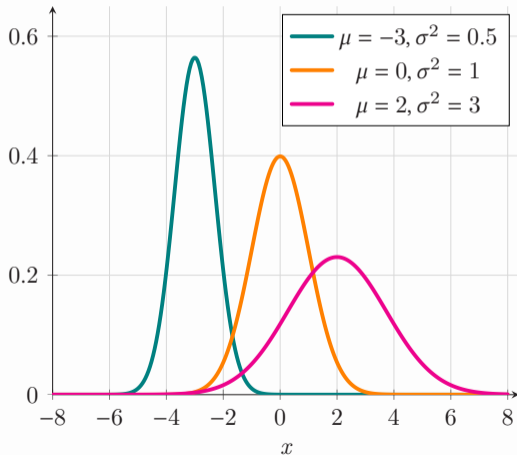
$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x - \mu)^2\right)$$

- There are two parameters, the **mean** μ which controls where the distribution is centred and the **variance** σ^2 which controls how wide it is

$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n \quad \hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \hat{\mu})^2$$

Parameters of the Univariate Gaussian Distribution

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- For binary classification, we begin by defining a model for each of our two classes
- We will make the *assumption* that, conditioned on the class, the data is Gaussian distributed

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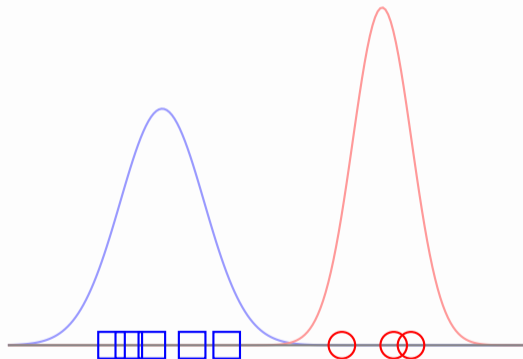
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- For data from **class 0**, we will assume that it is generated from $x|y = 0 \sim \mathcal{N}(x|\mu_0, \sigma_0^2)$
- For data from **class 1**, we will assume that it is generated from $x|y = 1 \sim \mathcal{N}(x|\mu_1, \sigma_1^2)$

Revisiting the 1D Example

- We can fit our two per-class Gaussians to our dataset

$$\mathcal{D} = \{(1.8, 0), (2.1, 0), (2.5, 0), (3.2, 0), (3.8, 0), (5.8, 1), (6.7, 1), (7.0, 1)\}$$



Generative Classifier - Making Predictions

- Now that we have a model for each class, and assuming that we have estimated the parameters for them (more on this later), we can use them to make predictions
- For a new test datapoint x we can simply assign it to the class with the *largest* output

$$\hat{y} = \arg \max_c \mathcal{N}(x | \mu_c, \sigma_c^2)$$

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- We may also want to know how ‘likely’ it is that a test datapoint is from a given class, e.g. from class 1

$$\hat{p}_1 = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)}{\mathcal{N}(x|\mu_0, \sigma_0^2) + \mathcal{N}(x|\mu_1, \sigma_1^2)}$$

where $\hat{p}_1 \in [0, 1]$



Adding 'Prior' Knowledge

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- We can encode this information as a weighting factor for each class, ϕ_0 and ϕ_1 , where $\phi_1, \phi_0 \in [0, 1]$
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- In the binary case $\phi_1 = 1 - \phi_0$, i.e. $\phi_0 + \phi_1 = 1$
- We can then combine this with the expression from the previous slide to obtain

$$\hat{p}_1 = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}{\mathcal{N}(x|\mu_0, \sigma_0^2)\phi_0 + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}$$

Bayes Classifier

- We came up with the following expression for making predictions for new data

$$\hat{p}_1 = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}{\mathcal{N}(x|\mu_0, \sigma_0^2)\phi_0 + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}$$

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- It turns out that this is just a restatement of **Bayes' rule**

$$p(y = c|x) = \frac{p(x|y = c)p(y = c)}{\sum_{c'} p(x|y = c')p(y = c')} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$

- Note, here we have omitted the dependence on the parameters for simplicity

Bayes' Rule

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- $p(y|x)$ is the **posterior** distribution of y , conditioned on x
- $p(x|y)$ is the **likelihood** of x , conditioned on y
- $p(y)$ is the **prior** distribution over y , i.e. what we know about y before seeing any data
- $p(x)$ is the **evidence**, which can be computed by marginalising over the unknown y , i.e. $\sum_y p(x|y)p(y)$

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IN HONOUR OF
THOMAS BAYES FRS

c. 1702 - 1761.

BAYES' THEOREM

$$P(X|Y) = \frac{P(Y|X) P(X)}{P(Y)}$$

Maximum Likelihood Estimation

Maximum Likelihood Estimation

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$$\mathcal{D} = \{(x_n, y_n)\}_{n=1}^{N_{\mathcal{D}}}$$
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$$\hat{\theta}_{\text{MLE}} = \arg \max_{\theta} p(\mathcal{D}|\theta)$$

Independence Assumption

- For convenience, we typically assume that the training data are *independent and identically* sampled from the same distribution, i.e. the **iid assumption**

$$p(\mathcal{D}|\boldsymbol{\theta}) = \prod_{n=1}^{N_{\mathcal{D}}} p(x_n, y_n; \boldsymbol{\theta})$$

Log Likelihood

- Taking the product of many terms can introduce numerical issues. To overcome this, we take the log which will not impact where the maximum of the function is

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- Recall that the log of a product equals the sum of the logs, i.e.
 $\log(ab) = \log(a) + \log(b)$

Negative Log Likelihood

- Many optimisation algorithms are designed to **minimise** functions. We can instead write the log likelihood (LL) as the **Negative Log Likelihood (NLL)**

$$\text{NLL}(\boldsymbol{\theta}) = - \sum_{n=1}^{N_{\mathcal{D}}} \log p(x_n, y_n; \boldsymbol{\theta})$$

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- Maximising the LL is equivalent to minimising the NLL

$$\hat{\boldsymbol{\theta}}_{\text{MLE}} = \arg \min_{\boldsymbol{\theta}} \text{NLL}(\boldsymbol{\theta})$$

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- These two terms depend on different sets of parameters $\boldsymbol{\theta} = \{\boldsymbol{\theta}_b, \boldsymbol{\theta}_g\}$, so they can be optimised independently

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- We can rewrite this as

$$\text{Ber}(y|\phi) = \phi^y (1 - \phi)^{(1-y)}$$

MLE for the Bernoulli Distribution

- We can compute the NLL for the Bernoulli with $\theta_b = \{\phi\}$ as follows

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- The MLE can be found by solving $\frac{\partial}{\partial \phi} \text{NLL}(\phi) = 0$
- Which results in

$$\hat{\phi} = \frac{N_1}{N_0 + N_1}$$

Gaussian Likelihood

- For the Gaussian NLL we need to solve for the parameters $\theta_g = \{\mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$, i.e. the parameters for both Gaussians (one for each class)

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Splitting the Data

- For convenience we will split the data into two subsets \mathcal{D}_0 and \mathcal{D}_1 , where $N_0 = |\mathcal{D}_0|$ and $N_1 = |\mathcal{D}_1|$
- Here, $\mathcal{D}_0 \subset \mathcal{D}$ is the subset of data where $y_n = 0$, and \mathcal{D}_1 is the subset where $y_n = 1$
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- Our expression for the Gaussian NLL now becomes

$$\text{NLL}(\theta_g) = - \sum_{x_n \in \mathcal{D}_0} \log \mathcal{N}(x_n | \mu_0, \sigma_0^2) - \sum_{x_n \in \mathcal{D}_1} \log \mathcal{N}(x_n | \mu_1, \sigma_1^2)$$

MLE for Univariate Gaussians

- Here, we will just focus on one of the Gaussians, i.e. the case where $y_n = 0$

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- The minimum of the NLL must satisfy the following conditions

$$\frac{\partial}{\partial \mu_0} \text{NLL}(\mu_0, \sigma_0^2) = 0, \quad \frac{\partial}{\partial \sigma_0^2} \text{NLL}(\mu_0, \sigma_0^2) = 0$$

MLE Solution for Univariate Gaussians

- Solving for the MLE for both classes we get the following expressions for the means

$$\hat{\mu}_0 = \frac{1}{N_0} \sum_{x_n \in \mathcal{D}_0} x_n, \quad \hat{\mu}_1 = \frac{1}{N_1} \sum_{x_n \in \mathcal{D}_1} x_n$$

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- With the following for the **variances**

$$\hat{\sigma}_0^2 = \frac{1}{N_0} \sum_{x_n \in \mathcal{D}_0} (x_n - \hat{\mu}_0)^2, \quad \hat{\sigma}_1^2 = \frac{1}{N_1} \sum_{x_n \in \mathcal{D}_1} (x_n - \hat{\mu}_1)^2$$

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- We have solved for the parameters $\theta = \{\phi, \mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$ of our model using MLE

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- We have solved for the parameters $\theta = \{\phi, \mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$ of our model using MLE
- Which we can use in our Bayes classifier

$$p(y = 1|x) = \frac{p(x|y = 1)p(y = 1)}{p(x|y = 0)p(y = 0) + p(x|y = 1)p(y = 1)}$$



Bringing it all Together

- We have solved for the parameters $\theta = \{\phi, \mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$ of our model using MLE
- Which we can use in our Bayes classifier

$$p(y = 1|x) = \frac{p(x|y = 1)p(y = 1)}{p(x|y = 0)p(y = 0) + p(x|y = 1)p(y = 1)}$$

- Which in the case of our binary classification model, is equivalent to

$$p(y = 1|x) = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi}{\mathcal{N}(x|\mu_0, \sigma_0^2)(1 - \phi) + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi}$$

Multivariate Classification

Multivariate Data

- Previously we discussed the case where the input feature was a one dimensional continuous value, i.e. $x \in \mathbb{R}$
- In practice, most datasets will be multivariate, i.e. $\mathbf{x} \in \mathbb{R}^D$
- We need to define model for multivariate data

Multivariate Gaussian

- The probability density function (PDF) of the **multivariate Gaussian** is given by

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{(D/2)}|\boldsymbol{\Sigma}|^{1/2}} \exp\left(-0.5(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)$$

- Here, $\boldsymbol{\mu} \in \mathbb{R}^D$ is the **mean vector** and $\boldsymbol{\Sigma} \in \mathbb{R}^{D \times D}$ is the **covariance matrix**
- The *univariate* Gaussian is a special case of this PDF



MLE for Multivariate Gaussian

- The maximum likelihood estimate of the **mean** vector is defined as

$$\hat{\boldsymbol{\mu}} = \frac{1}{N} \sum_{n=1}^N \mathbf{x}_n$$

MLE for Multivariate Gaussian

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- The maximum likelihood estimate of the **covariance matrix** is defined as

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \hat{\boldsymbol{\mu}})(\mathbf{x}_n - \hat{\boldsymbol{\mu}})^\top$$



Properties of the Covariance Matrix

- It is a square matrix ($D \times D$) specifying the covariance between each pair of elements of a given random vector
- Intuitively, it generalises the notion of variance to *multiple dimensions*
- The main diagonal contains variances, i.e. the covariance of each dimension with itself

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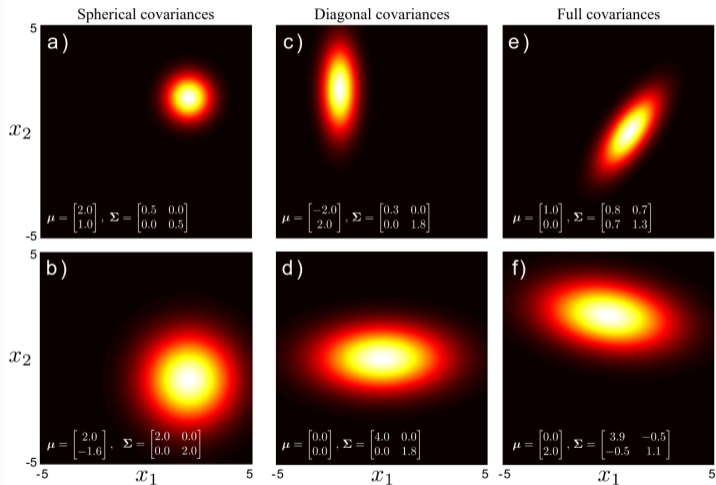
- The covariance matrix is **symmetric**, i.e. $\Sigma = \Sigma^T$ and $\Sigma^{-1} = (\Sigma^{-1})^T$
- It is positive semi-definite, i.e. $\mathbf{x}^T \Sigma \mathbf{x} \geq 0$ and $\mathbf{x}^T \Sigma^{-1} \mathbf{x} \geq 0$
- The full covariance matrix has $D(D+1)/2$ free parameters

Types of Covariance Matrices

- There are three types of covariance matrix
- Here, we show some 2D examples

$$\Sigma_{\text{spher}} = \begin{bmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{bmatrix} \quad \Sigma_{\text{diag}} = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \quad \Sigma_{\text{full}} = \begin{bmatrix} \sigma_{11}^2 & \sigma_{12}^2 \\ \sigma_{21}^2 & \sigma_{22}^2 \end{bmatrix}$$

Types of Covariance Matrices



Classification With Multivariate Gaussians

- We can use the same generative classification model as before

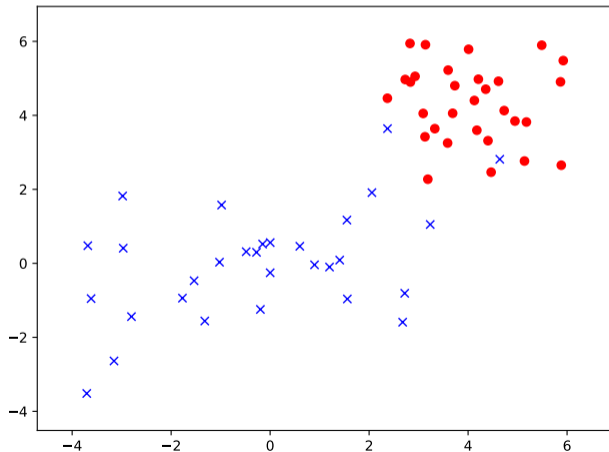
$$p(y = c|\mathbf{x}) = \frac{p(\mathbf{x}|y = c)p(y = c)}{\sum_{c'} p(\mathbf{x}|y = c')p(y = c')}$$

- In the multivariate case, we use a multivariate Gaussian for the class conditional density

$$p(\mathbf{x}|y = c) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c)$$

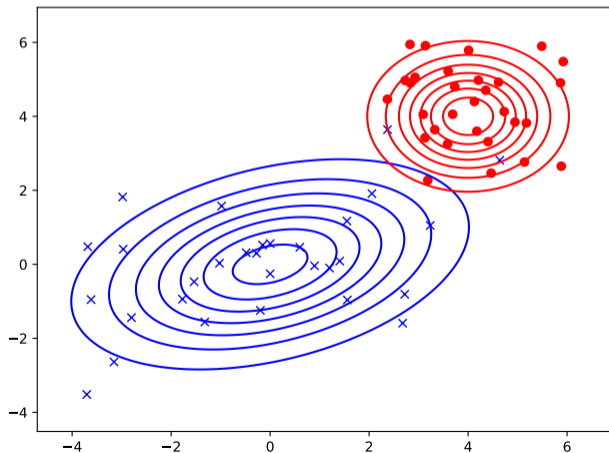
Gaussian Discriminant Analysis - 2D Example

- In this example we have two dimensional data from two different classes, **blue** and **red**



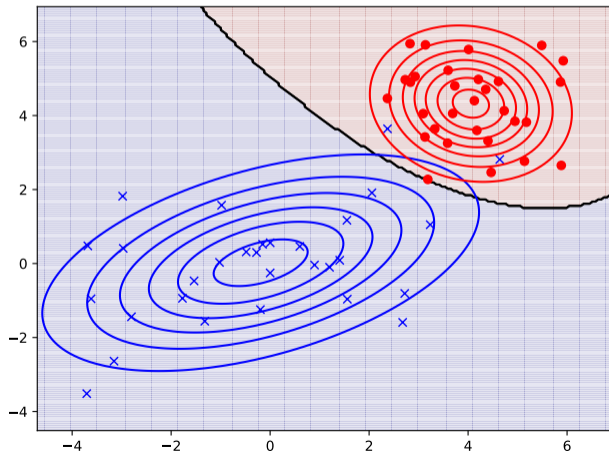
Gaussian Discriminant Analysis - 2D Example

- Here we visualise the underlying Gaussian distributions that generated the observed data



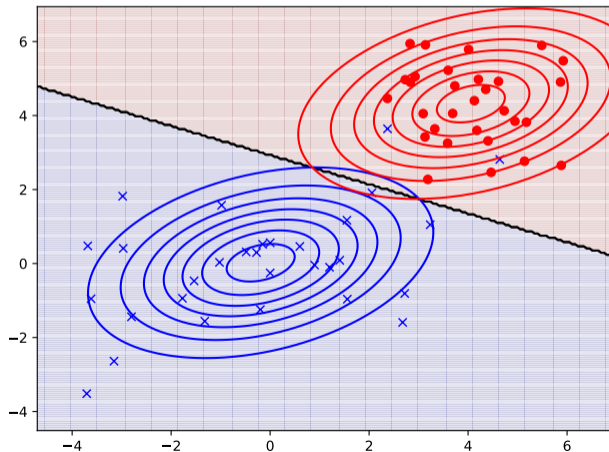
Quadratic Discriminant Analysis - 2D Example

- If we estimate a separate covariance matrix for each class (i.e. Σ_0 and Σ_1) and fit our classifier we get a **quadratic** decision boundary



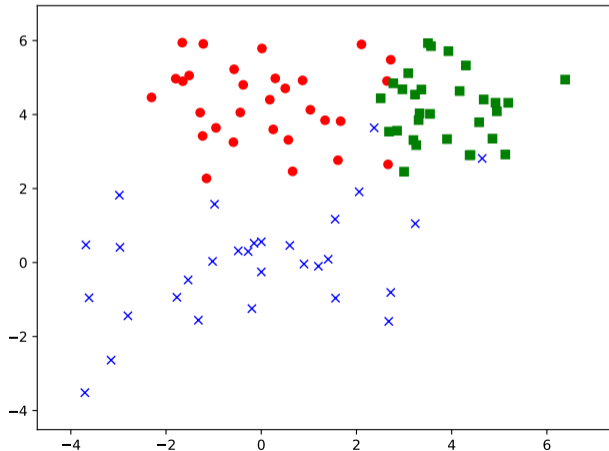
Linear Discriminant Analysis - 2D Example

- If instead, we assume that both classes share the same covariance matrix (i.e. $\Sigma_0 = \Sigma_1$) and fit our classifier we get a **linear** decision boundary



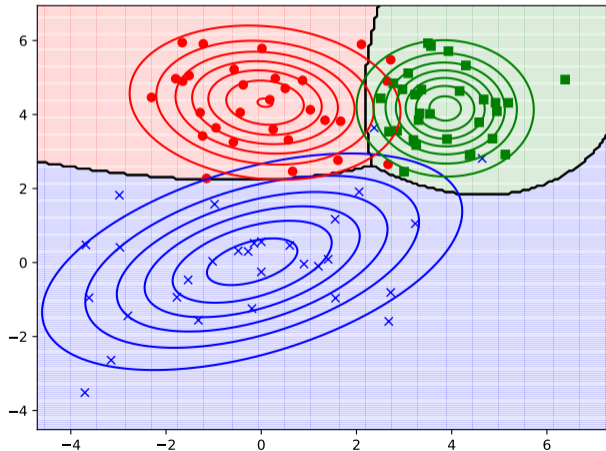
Multiclass Classification

- We can apply the same model in the multiclass case, i.e. where $y \in \{1, \dots, C\}$ and $C > 2$, by simply defining a class conditional model $p(\mathbf{x}|y = c)$ for each class



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Summary

- We introduced the problem of supervised classification
- We showed that simple Gaussian based models can be used for classification with continuous data through the application of Bayes' rule
- The parameters of these models are estimated using maximum likelihood estimation
- These models can be used for both single or vector input data and for binary or multiclass outputs