## **Applied Machine Learning (AML)**

Introduction to Classification

Oisin Mac Aodha • Siddharth N.

## **Classification Overview**

- ullet In supervised learning, we are tasked with predicting an output y, given an input feature vector  $m{x}$
- For **classification** problems, the output space is a set of mutually exclusive 'classes' (also commonly referred to as 'labels')

#### Classification

## **Binary versus Multiclass Classification**

- In **binary classification** we have two possibilities, e.g. dog versus cat. Thus,  $y \in \{0, 1\}, y \in \{1, 2\}, y \in \{-1, +1\}, \dots$
- In multiclass classification we can have C possible options, e.g. different breeds of dog. Thus,  $y \in \{1, ..., C\}$ , where C is the number of classes of interest

## **Example Classification Problems**

- Spam filtering
- Determining the object present in an image, i.e. image classification
- Fraudulent transaction detection
- Music genre classification
- Medical diagnostic tests

## informatics

3



## **Example 1D Classification Problem**

- We have collected a dataset consisting of the measurements of the petal length (in cm) of two different species of plants: species A and species B
- Thus, we have a one dimensional (1D) continuous measurement  $x \in \mathbb{R}$  and a binary class label  $y \in \{0, 1\}$
- $\bullet$  For species A, we have five measurements  $\{1.8, 2.1, 2.5, 3.2, 3.8\}$  and for species B we have three  $\{5.8, 6.7, 7.0\}$
- We can write our dataset  $\mathcal{D} = \{(x_n, y_n)\}_{n=1}^N =$  $\{(1.8,0), (2.1,0), (2.5,0), (3.2,0), (3.8,0), (5.8,1), (6.7,1), (7.0,1)\}$









## informatics

#### informatics

## **Example 1D Classification Problem**

- We have collected a dataset containing the measurements of the petal lengths (in cm) of plants from two different species: species A and species B
- Thus, we have a one dimensional (1D) continuous measurement  $x \in \mathbb{R}$  and a binary class label  $y \in \{0, 1\}$



## The Generative Approach

- Given a new observation x, can we predict which of the two classes it most likely belongs to?
- To do this, one approach is to fit a model to our already observed data
- We can then use this model to make predictions about unobserved (i.e. new) data
- For continuous features, one obvious choice is the Gaussian distribution

#### **Univariate Gaussian Distribution**

- The Gaussian (normal) distribution is a very widely used distribution for real-valued random variables, i.e.  $x \in \mathbb{R}$
- The probability density function of the Gaussian is defined as

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x-\mu)^2\right)$$

• There are two parameters, the **mean**  $\mu$  which controls where the distribution is centred and the **variance**  $\sigma^2$  which controls how wide it is

$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^{N} x_n \quad \hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^{N} (x_n - \hat{\mu})^2$$

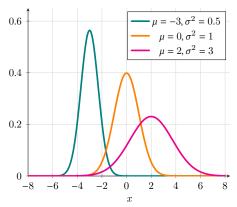


7

9

#### Parameters of the Univariate Gaussian Distribution

• The **mean**  $\mu$  controls where the distribution is centred and the **variance**  $\sigma^2$  controls how wide it is



#### informatics

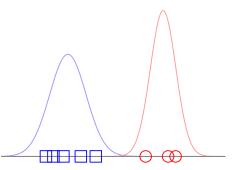
informatics

#### **Generative Classifier**

- For binary classification, we begin by defining a model for each of our two classes
- We will make the *assumption* that, conditioned on the class, the data is Gaussian distributed
- For data from class 0, we will assume that it is generated from  $x|y=0 \sim \mathcal{N}(x|\mu_0, \sigma_0^2)$
- For data from class 1, we will assume that it is generated from  $x|y=1 \sim \mathcal{N}(x|\mu_1,\sigma_1^2)$

## Revisiting the 1D Example

• We can fit our two per-class Gaussians to our dataset  $\mathcal{D} = \{(1.8,0), (2.1,0), (2.5,0), (3.2,0), (3.8,0), (5.8,1), (6.7,1), (7.0,1)\}$ 





## **Generative Classifier - Making Predictions**

- Now that we have a model for each class, and assuming that we have estimated the parameters for them (more on this later), we can use them to make predictions
- For a new test datapoint x we can simply assign it to the class with the *largest* output

$$\hat{y} = \arg\max_{c} \mathcal{N}(x|\mu_{c}, \sigma_{c}^{2})$$

• We may also want to know how 'likely' it is that a test datapoint is from a given class, e.g. from class 1

$$\hat{p_1} = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)}{\mathcal{N}(x|\mu_0, \sigma_0^2) + \mathcal{N}(x|\mu_1, \sigma_1^2)}$$

where  $\hat{p}_1 \in [0, 1]$ 



• In many cases, we made have prior knowledge that is relevant to our classification

For example, we may have many more observations from one class than another
We can encode this information as a weighting factor for each class, φ<sub>0</sub> and φ<sub>1</sub>, where

• We can then combine this with the expression from the previous slide to obtain

 $\hat{p_1} = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}{\mathcal{N}(x|\mu_0, \sigma_2^2)\phi_0 + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}$ 

## **Bayes Classifier**

We came up with the following expression for making predictions for new data

$$\hat{p_1} = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}{\mathcal{N}(x|\mu_0, \sigma_0^2)\phi_0 + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi_1}$$

• It turns out that this is just a restatement of **Bayes' rule** 

$$p(y = c|x) = \frac{p(x|y = c)p(y = c)}{\sum_{c'} p(x|y = c')p(y = c')} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$

• Note, here we have omitted the dependence on the parameters for simplicity

## Bayes' Rule

informatics

• Bayes' rule can be derived though application of the *product rule*, i.e.

$$p(x, y) = p(x|y)p(y) = p(y|x)p(x)$$

Adding 'Prior' Knowledge

• In the binary case  $\phi_1 = 1 - \phi_0$ , i.e.  $\phi_0 + \phi_1 = 1$ 

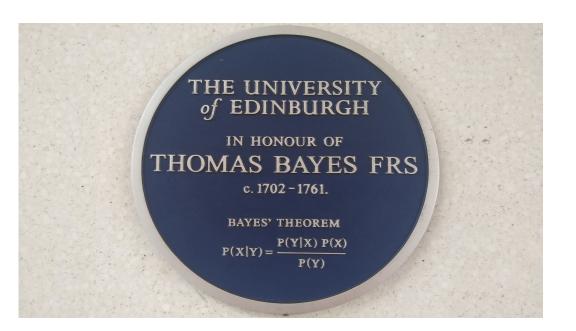
problem

 $\phi_1, \phi_0 \in [0, 1]$ 

$$p(y|x) = \frac{p(x|y)p(y)}{p(x)}$$

- p(y|x) is the **posterior** distribution of y, conditioned on x
- p(x|y) is the **likelihood** of x, conditioned on y
- ullet p(y) is the **prior** distribution over y, i.e. what we know about y before seeing any data
- p(x) is the **evidence**, which can be computed by marginalising over the unknown y, i.e.  $\sum_y p(x|y)p(y)$

12



#### **Maximum Likelihood Estimation**

## **Maximum Likelihood Estimation**

- In binary classification we have a set of  $N_{\mathcal{D}}$  pairs of observations, where  $\mathcal{D} = \{(x_n, y_n)\}_{n=1}^{N_{\mathcal{D}}}$
- The process of learning the model parameters  $\theta$  from our dataset  $\mathcal D$  is called **model fitting** or **training**
- One common approach for fitting a model to data, is called Maximum Likelihood Estimation (MLE)
- Here we aim to find the parameters that assign the highest *likelihood* to our data given our model, i.e. the ones that maximise the likelihood

$$\hat{\boldsymbol{\theta}}_{\mathsf{MLE}} = \mathop{\arg\max}_{\boldsymbol{\theta}} p(\mathcal{D}|\boldsymbol{\theta})$$

## **Independence Assumption**

• For convenience, we typically assume that the training data are *independent and identically* sampled from the same distribution, i.e. the **iid assumption** 

$$p(\mathcal{D}|\boldsymbol{\theta}) = \prod_{n=1}^{N_{\mathcal{D}}} p(x_n, y_n; \boldsymbol{\theta})$$



## Log Likelihood

• Taking the product of many terms can introduce numerical issues. To overcome this, we take the log which will not impact where the maximum of the function is

$$\begin{aligned} \mathsf{LL}(\boldsymbol{\theta}) &= \log p(\mathcal{D}|\boldsymbol{\theta}) \\ &= \log \prod_{n=1}^{N_{\mathcal{D}}} p(x_n, y_n; \boldsymbol{\theta}) \\ &= \sum_{n=1}^{N_{\mathcal{D}}} \log p(x_n, y_n; \boldsymbol{\theta}) \end{aligned}$$

• Recall that the log of a product equals the sum of the logs, i.e. log(ab) = log(a) + log(b)



18

informatics

## **Negative Log Likelihood**

 Many optimisation algorithms are designed to minimise functions. We can instead write the log likelihood (LL) as the Negative Log Likelihood (NLL)

$$\mathsf{NLL}(\boldsymbol{\theta}) = -\sum_{n=1}^{N_{\mathcal{D}}} \log p(x_n, y_n; \boldsymbol{\theta})$$

• Maximising the LL is equivalent to minimising the NLL

$$\hat{\theta}_{\mathsf{MLE}} = \operatorname*{arg\,min}_{\boldsymbol{\theta}} \mathsf{NLL}(\boldsymbol{\theta})$$

## **Negative Log Likelihood**

We can rewrite our expression for the NLL as

$$\begin{aligned} \text{NLL}(\boldsymbol{\theta}) &= -\sum_{n=1}^{N_{\mathcal{D}}} \log p(x_n, y_n; \boldsymbol{\theta}) \\ &= -\sum_{n=1}^{N_{\mathcal{D}}} \log \left[ p(y_n; \boldsymbol{\theta}_b) p(x_n | y_n; \boldsymbol{\theta}_g) \right] \\ &= -\left[ \sum_{n=1}^{N_{\mathcal{D}}} \log p(y_n; \boldsymbol{\theta}_b) \right] - \left[ \sum_{n=1}^{N_{\mathcal{D}}} \log p(x_n | y_n; \boldsymbol{\theta}_g) \right] \end{aligned}$$
Bernoulli NLL of labels
Guassian NLL of features

• These two terms depend on different sets of parameters  $\theta = \{\theta_b, \theta_g\}$ , so they can be optimised independently

#### **Bernoulli Distribution**

- In the case of the binary label data  $y \in \{0, 1\}$ , we can use a Bernoulli prior
- $\bullet\;$  The probability mass function with the parameter  $\phi$  of the Bernoulli is defined as

$$Ber(y|\phi) = \begin{cases} 1 - \phi & \text{if } y = 0\\ \phi & \text{if } y = 1 \end{cases}$$

• We can rewrite this as

Ber
$$(y|\phi) = \phi^y (1 - \phi)^{(1-y)}$$

#### MLE for the Bernoulli Distribution

• We can compute the NLL for the Bernoulli with  $\theta_b = \{\phi\}$  as follows

$$\begin{aligned} \mathsf{NLL}(\phi) &= -\sum_{n=1}^{N_{\mathcal{D}}} \log p(y_n; \boldsymbol{\theta}_b) \\ &= -\sum_{n=1}^{N_{\mathcal{D}}} \log \left[ \phi^{y_n} (1 - \phi)^{(1 - y_n)} \right] \\ &= -N_1 \log(\phi) - N_0 \log(1 - \phi) \end{aligned}$$

- The MLE can be found by solving  $\frac{\partial}{\partial \phi} NLL(\phi) = 0$
- Which results in

$$\hat{\phi} = \frac{N_1}{N_0 + N_1}$$

#### informatics

## 22 THE UNIVERSITY OF EDINBURGH INFORMATICS

 $= -\sum_{n=0}^{N_{\mathcal{B}}} \log \left[ \mathcal{N}(x_n | \mu_0, \sigma_0^2)^{(1-y_n)} \mathcal{N}(x_n | \mu_1, \sigma_1^2)^{(y_n)} \right]$ 

 $= -\sum_{n=1}^{N_{\mathcal{D}}} (1 - y_n) \log \left[ \mathcal{N}(x_n | \mu_0, \sigma_0^2) \right] - \sum_{n=1}^{N_{\mathcal{D}}} y_n \log \left[ \mathcal{N}(x_n | \mu_1, \sigma_1^2) \right]$ 

• For the Gaussian NLL we need to solve for the parameters  $\theta_q = \{\mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$ , i.e. the

parameters for both Gaussians (one for each class)

 $NLL(\mu_0, \sigma_0^2, \mu_1, \sigma_1^2) = -\sum_{n=0}^{N_D} \log p(x_n | y_n; \boldsymbol{\theta}_g)$ 

## Splitting the Data

- For convenience we will split the data into two subsets  $\mathcal{D}_0$  and  $\mathcal{D}_1$ , where  $N_0 = |\mathcal{D}_0|$  and  $N_1 = |\mathcal{D}_1|$
- Here,  $\mathcal{D}_0 \subset \mathcal{D}$  is the subset of data where  $y_n = 0$ , and  $\mathcal{D}_1$  is the subset where  $y_n = 1$
- We can then find the maximum likelihood estimate for each set separately
- Our expression for the Guassian NLL now becomes

$$\mathsf{NLL}(\boldsymbol{\theta}_g) = -\sum_{x_n \in \mathcal{D}_0} \log \mathcal{N}(x_n | \mu_0, \sigma_0^2) - \sum_{x_n \in \mathcal{D}_1} \log \mathcal{N}(x_n | \mu_1, \sigma_1^2)$$

#### **MLE for Univariate Gaussians**

Gaussian Likelihood

ullet Here, we will just focus on one of the Gaussians, i.e. the case where  $y_n=0$ 

$$\begin{aligned} \mathsf{NLL}(\mu_0, \sigma_0^2) &= -\sum_{x_n \in \mathcal{D}_0} \log \mathcal{N}(x_n | \mu_0, \sigma_0^2) \\ &= -\sum_{x_n \in \mathcal{D}_0} \log \left[ \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp\left( -\frac{1}{2\sigma_0^2} (x_n - \mu_0)^2 \right) \right] \\ &= \frac{N_0}{2} \log(2\pi) + \frac{N_0}{2} \log(\sigma_0^2) + \sum_{x_n \in \mathcal{D}_0} \frac{(x_n - \mu_0)^2}{2\sigma_0^2} \end{aligned}$$

• The minimum of the NLL must satisfy the following conditions

$$\frac{\partial}{\partial \mu_0} \text{NLL}(\mu_0, \sigma_0^2) = 0, \qquad \frac{\partial}{\partial \sigma_0^2} \text{NLL}(\mu_0, \sigma_0^2) = 0$$

24

## **MLE Solution for Univariate Gaussians**

• Solving for the MLE for both classes we get the following expressions for the means

$$\hat{\mu_0} = \frac{1}{N_0} \sum_{x_n \in \mathcal{D}_0} x_n, \qquad \hat{\mu_1} = \frac{1}{N_1} \sum_{x_n \in \mathcal{D}_1} x_n$$

• With the following for the variances

$$\hat{\sigma_0}^2 = \frac{1}{N_0} \sum_{x_n \in \mathcal{D}_0} (x_n - \hat{\mu_0})^2, \qquad \hat{\sigma_1}^2 = \frac{1}{N_1} \sum_{x_n \in \mathcal{D}_1} (x_n - \hat{\mu_1})^2$$



#### **Multivariate Classification**

## Bringing it all Together

- We have solved for the parameters  $\theta = \{\phi, \mu_0, \sigma_0^2, \mu_1, \sigma_1^2\}$  of our model using MLE
- Which we can use in our Bayes classifier

$$p(y=1|x) = \frac{p(x|y=1)p(y=1)}{p(x|y=0)p(y=0) + p(x|y=1)p(y=1)}$$

• Which in the case of our binary classification model, is equivalent to

$$p(y=1|x) = \frac{\mathcal{N}(x|\mu_1, \sigma_1^2)\phi}{\mathcal{N}(x|\mu_0, \sigma_0^2)(1-\phi) + \mathcal{N}(x|\mu_1, \sigma_1^2)\phi}$$

# informatics

26

#### **Multivariate Data**

- Previously we discussed the case where the input feature was a one dimensional continuous value, i.e.  $x \in \mathbb{R}$
- ullet In practice, most datasets will be multivariate, i.e.  $oldsymbol{x} \in \mathbb{R}^D$
- We need to define model for multivariate data



#### **Multivariate Gaussian**

• The probability density function (PDF) of the multivariate Gaussian is given by

$$\mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu},\boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{(D/2)}|\boldsymbol{\Sigma}|^{1/2}} \exp\left(-0.5(\boldsymbol{x}-\boldsymbol{\mu})^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}(\boldsymbol{x}-\boldsymbol{\mu})\right)$$

- Here,  $\mu \in \mathbb{R}^D$  is the mean vector and  $\Sigma \in \mathbb{R}^{D \times D}$  is the covariance matrix
- The univariate Gaussian is a special case of this PDF

#### informatics

29

31



# Properties of the Covariance Matrix

- It is a square matrix  $(D \times D)$  specifying the covariance between each pair of elements of a given random vector
- Intuitively, it generalises the notion of variance to *multiple dimensions*
- The main diagonal contains variances, i.e. the covariance of each dimension with itself
- The covariance matrix is **symmetric**, i.e.  $\Sigma = \Sigma^{T}$  and  $\Sigma^{-1} = (\Sigma^{-1})^{T}$
- It is positive semi-definite, i.e.  $x^{\mathsf{T}} \Sigma x \geq 0$  and  $x^{\mathsf{T}} \Sigma^{-1} x \geq 0$
- The full covariance matric has D(D+1)/2 free parameters

#### **MLE for Multivariate Gaussian**

• The maximum likelihood estimate of the mean vector is defined as

$$\hat{\boldsymbol{\mu}} = \frac{1}{N} \sum_{n=1}^{N} \boldsymbol{x}_n$$

• The maximum likelihood estimate of the **covariance matrix** is defined as

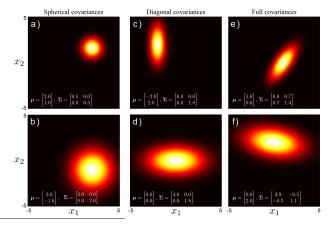
$$\hat{\Sigma} = \frac{1}{N} \sum_{n=1}^{N} (\boldsymbol{x}_n - \hat{\boldsymbol{\mu}}) (\boldsymbol{x}_n - \hat{\boldsymbol{\mu}})^{\mathsf{T}}$$

## **Types of Covariance Matrices**

- There are three types of covariance matrix
- Here, we show some 2D examples

$$\Sigma_{\mathsf{spher}} = \begin{bmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{bmatrix} \quad \Sigma_{\mathsf{diag}} = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \quad \Sigma_{\mathsf{full}} = \begin{bmatrix} \sigma_{11}^2 & \sigma_{12}^2 \\ \sigma_{21}^2 & \sigma_{22}^2 \end{bmatrix}$$

## **Types of Covariance Matrices**



Simon Prince - Computer Vision Models (Book)

informatics

#### **Classification With Multivariate Gaussians**

• We can use the same generative classification model as before

$$p(y = c|\mathbf{x}) = \frac{p(\mathbf{x}|y = c)p(y = c)}{\sum_{c'} p(\mathbf{x}|y = c')p(y = c')}$$

• In the multivariate case, we use a multivariate Gaussian for the class conditional density

$$p(\boldsymbol{x}|y=c) = \mathcal{N}(\boldsymbol{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c)$$

informatics

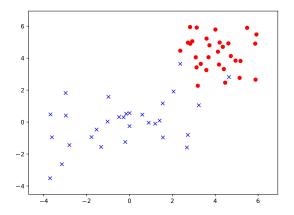
33

35

#### 34

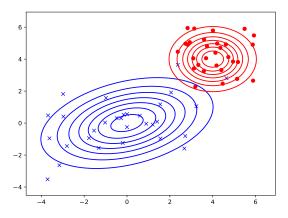
## Gaussian Discriminant Analysis - 2D Example

 In this example we have two dimensional data from two different classes, blue and red



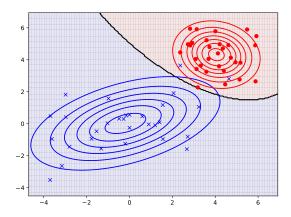
## Gaussian Discriminant Analysis - 2D Example

• Here we visualise the underlying Gaussian distributions that generated the observed data



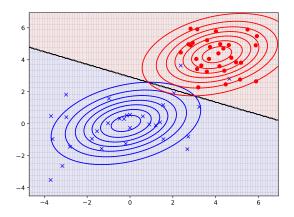
## Quadratic Discriminant Analysis - 2D Example

• If we estimate a separate covariance matrix for each class (i.e.  $\Sigma_0$  and  $\Sigma_1$ ) and fit our classifier we get a **quadratic** decision boundary



Linear Discriminant Analysis - 2D Example

• If instead, we assume that both classes share the same covariance matrix (i.e.  $\Sigma_0 = \Sigma_1$ ) and fit our classifier we get a **linear** decision boundary

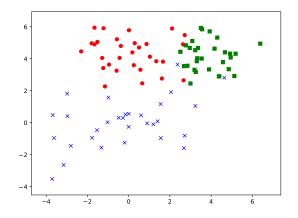


THE UNIVERSITY & EDINBURGH INFORMATICS

#### THE UNIVERSITY of EDINBURGH INFORMATICS

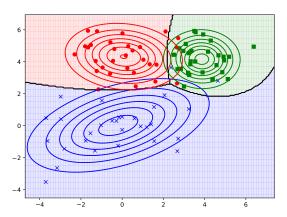
#### **Multiclass Classification**

• We can apply the same model in the multiclass case, i.e. where  $y \in \{1, ..., C\}$  and C > 2, by simply defining a class conditional model p(x|y=c) for each class



#### **Multiclass Classification**

• We can apply the same model in the multiclass case, i.e. where  $y \in \{1,...,C\}$  and C > 2, by simply defining a class conditional model p(x|y=c) for each class



## **Summary**

- We introduced the problem of supervised classification
- We showed that simple Guassian based models can be used for classification with continuous data through the application of Bayes' rule
- The parameters of these models are estimated using maximum likelihood estimation
- These models can be used for both single or vector input data and for binary or multiclass outputs

